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## **► To cite this version:**

Mohamed El Hedi Arouri, Duc Khuong Nguyen. Oil Prices, Stock Markets and Portfolio Investment: Evidence from Sector Analysis in Europe over the Last Decade. 2010. hal-00507823

**HAL Id: hal-00507823**

**<https://hal.science/hal-00507823>**

Preprint submitted on 1 Aug 2010

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# Oil Prices, Stock Markets and Portfolio Investment: Evidence from Sector Analysis in Europe over the Last Decade\*

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## **Abstract**

This article extends the understanding of oil–stock market relationships over the last turbulent decade. Unlike previous empirical investigations, which have largely focused on broad-based market indices (national and/or regional indices), we examine short-term linkages in the aggregate as well as sector by sector levels in Europe using different econometric techniques. Our main findings suggest that the reactions of stock returns to oil price changes differ greatly depending on the activity sector. In the out-of-sample analysis we show that introducing oil asset into a diversified portfolio of stocks allows to significantly improve its risk-return characteristics.

**Keywords:** oil prices, portfolio management, short-term analysis, sector indices.

**JEL classifications:** G12, F3, Q43.

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\* The authors thank anonymous reviewers and Stéphane Grégoir for helpful comments and suggestions on an earlier version of the paper. The usual disclaimer applies.

## 1. Introduction

Understanding the dynamics of stock returns is an issue of ongoing research in financial market literature. In particular, identifying the factors that drive stock market returns is of utmost relevance and importance to investors and policy makers. Although an abundance of theoretical and empirical works focused on asset pricing, there is no consensus about both the nature and number of factors of stock returns. Furthermore, as oil price has changed with sequences of very large increases and decreases over recent years, it is now quite opportune to augment the existing research on its impacts on stock market returns. Various transmission channels exist through which oil price fluctuations may affect stock returns. Indeed, the value of stock in theory equals discounted sum of expected future cash-flows. These discounted cash-flows reflect economic conditions (e.g., inflation, interest rates, production costs, income, economic growth, and investor and consumer confidence) and macroeconomic events that are likely to be influenced by oil shocks. Accordingly, oil price changes may impact stock returns.

In the literature, there has been a large volume of works on the linkages between oil prices and economic variables. The majority of these studies have shown significant effects of oil price fluctuations on economic activity for several developed and emerging countries (Hamilton, 2003; Cunado and Perez de Garcia, 2005; Balaz and Londarev, 2006; Lardic and Mignon, 2008; Gronwald, 2008; Cologni and Manera, 2008; and Kilian, 2008). By contrast, there have been relatively a few attempts to study the dynamic relationship between oil price variations and stock markets.

The pioneering paper by Jones and Kaul (1996) tests the reaction of stock returns in four developed markets (Canada, Japan, the UK, and the US) to oil price fluctuations on the basis of the standard cash-flow dividend valuation model. They find that for the US and Canada stock market reaction can be accounted for entirely by the impact of oil shocks on cash flows. The results for Japan and the UK were nevertheless inconclusive. Using an unrestricted vector autoregressive (VAR) model, Huang *et al.* (1996) find no evidence of a relationship between oil prices and the S&P500 market index. Inversely, Sadorsky (1999) also applies an unrestricted VAR model with GARCH effects to American monthly data and shows a significant relationship between oil price changes and US aggregate stock returns. Park and Ratti (2008) show that oil price increases have a negative impact on stock returns in the US and twelve European countries, whereas stock markets in Norway, an oil-exporting country respond positively to rises of oil price. In a more recent study, Apergis and Miller (2009) also examine whether structural oil-market shocks affect stock returns in eight developed countries,

and document no significant responses of international stock market returns to oil price shocks.

Very few studies have looked at the impact of oil price changes on the stocks of individual sectors. In addition, most of these studies are country-specific and therefore do not provide a global perspective. For instance, Sadorsky (2001) and Boyer and Filion (2007) show that oil price increases lead to higher stock returns of Canadian oil and gas companies. El-Sharif *et al.* (2005) reach the same conclusion when analyzing oil and gas returns in the UK. However, the authors note that non-oil and gas sectors are weakly linked to oil price changes. More recently, Nandha and Faff (2008) study the short-term link between oil prices and thirty-five Datastream global industries and report that oil price rises have a negative impact on all, but not the oil and gas industries. Finally, Nandha and Brooks (2009) look into the reaction of the transport sector to oil prices in thirty-eight countries and find that, in developed economies, oil prices have some influence on the returns of the sector under consideration. There is however no evidence of a significant role for oil price changes in Asian and Latin American countries. Taken together, the results from the available works on the relationships between oil price changes and sector stock returns are inconclusive and differ from country to country.

The current article extends the understanding of the relationship between oil price changes and stock returns at the disaggregated sector level in Europe by investigating their short-term linkages over the last turbulent decade using different econometric techniques. Over this decade of globally increasing oil prices, the responses of stock markets to oil price changes are ambiguous. Indeed, on the one hand increases in oil prices translate into higher transportation, production, and heating costs, which can put a drag on corporate earnings. Rising oil prices can also stir up concerns about inflation and curtail consumers' discretionary spending. On the other hand, investors can also associate increasing oil prices with a booming economy. Thus, higher oil prices could reflect stronger business performance.

It is equally important to note that studying the short-term effects of oil price fluctuations at sector level instead of aggregate market level is important for several reasons. First, any market-wide consequence may hide the performance, not necessarily uniform, of various economic sectors. Further, sector sensitivities to changes in oil price can be asymmetric to the extent that some sectors may be more severely affected by these changes than the others. The degree to which a sector is more or less sensitive to oil depend upon whether oil serves as its input or output, its exposure to the indirect oil effects, its degree of competition and concentration, and its capacity to absorb and transfer oil price risk to its consumers. Second, the industrial base varies from one European market to another. Large and mature markets

such as France and Germany are more diversified, whereas small markets such as Switzerland usually concentrate on a few industries. Thus, the results of studies based on national stock market indices such as Park and Ratti (2008) and Apergis and Miller (2009) should be considered with precaution. An important and interesting issue consists then of examining how different sector market indices rather than national market indices react to oil price fluctuations. Finally, indentifying the heterogeneity of sector sensitivities to oil has important implications for portfolio risk management since some sectors may still provide a meaningful channel for international diversification during large swings in oil prices.

The rest of the paper is organized as follows. Section 2 presents the data and some preliminary analysis. Section 3 reports and discusses the empirical results. Section 4 focuses on some out-of-sample forecasting evaluations and portfolio implications of empirical results. Summary conclusions are provided in Section 5.

## **2. Data and preliminary analysis**

We investigate the relationships between oil prices and stock returns in Europe from a sector perspective. Our sample data include the Dow Jones (DJ) Stoxx 600 and twelve European sector indices, namely *Automobile & Parts*, *Financials*, *Food & Beverages*, *Oil & Gas*, *Health Care*, *Industrials*, *Basic Materials*, *Personal & Household Goods*, *Consumer Services*, *Technology*, *Telecommunications* and *Utilities*. We collect stock market data from Datastream database.

Introduced in 1998, the Dow Jones Stoxx 600 sector indices aim to represent the largest European companies in each of the most important industries and currently cover Austria, Belgium, Denmark, Finland, France, Germany, Greece, Iceland, Ireland, Italy, Luxembourg, the Netherlands, Norway, Portugal, Spain, Sweden, Switzerland and the United Kingdom. The sector indices offer an alternative view of the performance of the European stock markets. The main industries are Automobile and Parts (automobiles, auto parts and tires), Financials (banks, insurance, reinsurance, real estate and financial services), Food and Beverages (beverages and food producers), Oil and Gas (oil and gas producers, oil equipment, and services, distribution and alternative energy), Health Care (health care equipment and services, and pharmaceuticals and biotechnology), Industrials (construction and materials, and industrial goods and services), Basic Materials (chemicals and basic resources), Personal and Household Goods (household goods, home construction, leisure goods, and personal goods and tobacco), Consumer Services (retail, media, travel and leisure), Technology (software and computer services, and technology hardware and equipment), Telecommunications (fixed line

and mobile telecommunications), and Utilities (electricity, gas, water and multi-utilities). Each sector index represents a capitalization-weighted portfolio of the largest European companies in this sector.

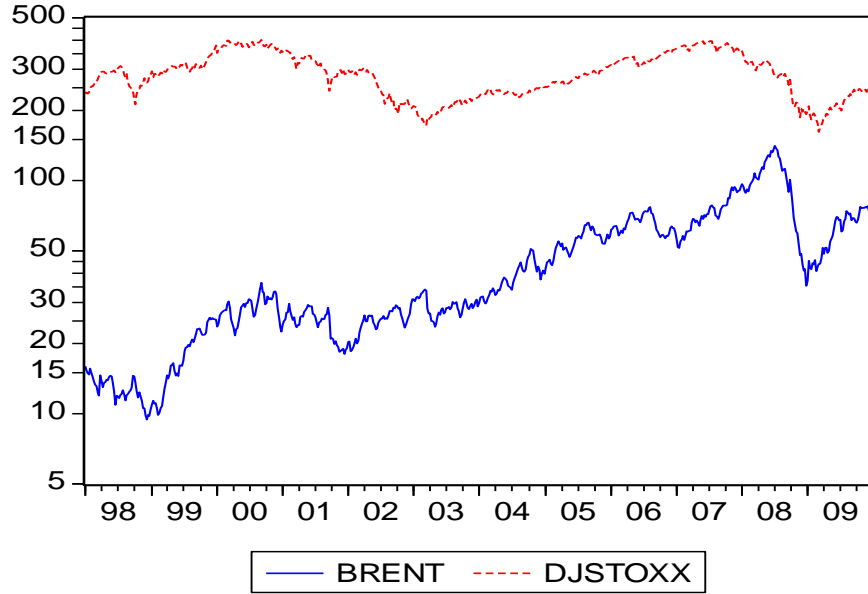
We think that weekly data may better capture the interaction of oil and stock price changes than daily or monthly data. On the one hand, the use of weekly data in the analysis instead of daily data significantly reduces any potential biases that may arise such as the bid-ask effect, non-synchronous trading days, etc. On the other hand, the monthly data may have some bearing on asymmetry in responses of stock returns to oil price shocks. In this schema of thing, we make use of weekly stock market sector indices over the period from January 01, 1998 to November 13, 2008 and examine their sensitivity to the recent oil price boom after the 1997 Asian financial crisis.<sup>1</sup> Over this sample period, the relationship between oil prices and stock markets was ambiguous as shown by Figure 1. Increases in oil prices were, at the same time, indicative of higher production costs and inflation pressure, and synonyms of higher expected economic growth and higher levels of consumer and investor confidence. Notice that weekly data running from November 20, 2008 to December 31, 2009 will be employed in our out-of-sample analysis to shed light on forecasting evaluation and some portfolio investment implications of the in-sample results.

For oil, we use the weekly Brent crude oil price obtained from the Energy Information Administration (EIA). The Europe Brent is one of the major international oil benchmarks. We express Brent oil prices in euro using euro/dollar exchange rates from Datastream.

**Figure 1. European market index (DJ Stoxx 600) and crude oil price (Brent)**

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<sup>1</sup> It should be noted that both daily and monthly data as well as longer sample period are employed to subsequently check the robustness of our results.



Before we can conduct further analysis on oil and stock market sector indices, the order of integration of our series is investigated using three standard unit root tests: Augmented Dickey-Fuller (ADF), Phillips-Perron (PP), and Kwiatkowski *et al.* (KPSS) tests. The ADF and PP tests are based on the null hypothesis of a unit root, while the KPSS test considers the null of no unit root. The obtained results are reported in Table 1. All the price series appear to be integrated of order one, which is a standard result in the literature for such series.

**Table 1. Unit root tests**

	Levels			First difference		
	ADF	PP	KPSS	ADF	PP	KPSS
Oil Brent	0.60 <sup>a</sup>	-2.68 <sup>b</sup>	0.69 <sup>**b</sup>	-11.48 <sup>*a</sup>	-18.35 <sup>*a</sup>	0.08 <sup>b</sup>
DJ Stoxx	-0.38 <sup>a</sup>	-0.26 <sup>a</sup>	0.37 <sup>*b</sup>	-9.16 <sup>*a</sup>	-23.6 <sup>*a</sup>	0.27 <sup>b</sup>
Automobile & Parts	0.03 <sup>a</sup>	-2.14 <sup>b</sup>	0.67 <sup>**b</sup>	-27.48 <sup>*a</sup>	-27.46 <sup>*a</sup>	0.058 <sup>b</sup>
Financials	-0.78 <sup>a</sup>	-0.66 <sup>c</sup>	0.26 <sup>b</sup>	-8.29 <sup>*a</sup>	-23.91 <sup>*c</sup>	0.35 <sup>*b</sup>
Food & Beverages	0.43 <sup>a</sup>	-2.11 <sup>b</sup>	1.55 <sup>*b</sup>	-21.83 <sup>*a</sup>	21.83 <sup>*a</sup>	0.08 <sup>b</sup>
Oil & Gas	0.09 <sup>a</sup>	-2.01 <sup>b</sup>	1.14 <sup>*b</sup>	-15.54 <sup>*a</sup>	-23.38 <sup>*a</sup>	0.14 <sup>b</sup>
Health Care	0.02 <sup>a</sup>	-2.07 <sup>b</sup>	0.22 <sup>b</sup>	-25.85 <sup>*a</sup>	-25.87 <sup>*a</sup>	0.12 <sup>b</sup>
Industrials	-0.48 <sup>a</sup>	-0.25 <sup>a</sup>	0.49 <sup>*b</sup>	-5.83 <sup>*a</sup>	-22.12 <sup>*a</sup>	0.19 <sup>b</sup>
Basic Materials	0.34 <sup>a</sup>	0.32 <sup>a</sup>	2.03 <sup>*b</sup>	-22.29 <sup>*a</sup>	-22.47 <sup>*a</sup>	0.14 <sup>b</sup>
Personal & Household Goods	0.30 <sup>a</sup>	1.87 <sup>a</sup>	1.55 <sup>*b</sup>	-23.19 <sup>*a</sup>	-23.23 <sup>*a</sup>	0.18 <sup>b</sup>
Consumer Services	-0.91 <sup>a</sup>	1.34 <sup>c</sup>	1.00 <sup>*b</sup>	-6.21 <sup>*a</sup>	-22.76 <sup>*a</sup>	0.23 <sup>b</sup>
Technology	-0.66 <sup>a</sup>	2.00 <sup>b</sup>	1.25 <sup>*b</sup>	-6.49 <sup>*a</sup>	-24.07 <sup>*a</sup>	0.21 <sup>b</sup>
Telecommunications	-0.30 <sup>a</sup>	-2.10 <sup>c</sup>	1.09 <sup>*b</sup>	-7.79 <sup>*a</sup>	-23.89 <sup>*a</sup>	0.22 <sup>b</sup>
Utilities	0.75 <sup>a</sup>	0.76 <sup>a</sup>	1.45 <sup>*b</sup>	17.4 <sup>*a</sup>	22.78 <sup>*a</sup>	0.19 <sup>b</sup>

Notes: All variables are in natural logs. ADF is the Augmented Dickey-Fuller test, PP the Phillips-Perron test, and KPSS the Kwiatkowski-Phillips-Schmidt-Shin test. (<sup>a</sup>) indicates a model without constant or deterministic trend, (<sup>b</sup>) a model with constant without deterministic trend, and (<sup>c</sup>) a model with constant and deterministic trend. \*, \*\* and \*\*\* denote rejection of the null hypothesis at the 1%, 5% and 10% levels respectively.

Descriptive statistics for return series (first logarithmic differences) are summarized in Table 2. On average, oil experienced higher returns than European stock market returns over our sample period. *Technology* stocks have the highest volatility followed by oil and *Automobile* stocks. Skewness is negative in most cases and the Jarque-Bera test statistic (JB)

strongly rejects the hypothesis of normality. There is also strong evidence of ARCH effects and there are significant serial correlations for some series.

**Table 2. Descriptive statistics of return series**

	Mean	Std. dev.	Skew.	Kurt.	ARCH test	LB	LB <sup>2</sup>	JB	Corr. with oil	Corr. with DJ Stoxx
Oil Brent	0.21	4.61	-0.47	4.72	23.81*	38.07*	29.63*	91.4*	1.00	0.10
DJ Stoxx	-0.02	2.56	-0.68	6.21	79.08*	10.62***	116.36*	291.9*	0.10	1.00
Automobile and Parts	0.02	4.32	0.54	18.80	127.28*	11.81***	143.77*	5930.1*	-0.02	0.67
Financials	-0.09	3.29	-0.53	7.53	118.60*	8.99	231.09*	512.7*	0.04	0.92
Food and Beverages	0.05	2.18	-0.54	5.82	27.27*	7.45	36.10*	217.0*	-0.10	0.57
Oil and Gas	0.01	3.18	-0.52	4.66	101.88*	15.16**	187.45*	91.3*	0.33	0.63
Health Care	0.01	2.54	-0.03	5.23	44.93*	9.72	61.04*	118.5*	-0.07	0.63
Industrials	-0.02	2.90	-0.71	6.52	81.71*	12.68**	98.56*	341.6*	0.10	0.91
Basic Materials	0.06	3.34	-0.90	7.02	140.87*	7.92	230.38*	459.5*	0.12	0.76
Personal and Household Goods	0.04	2.57	-0.76	6.70	40.53*	6.94	50.67*	378.6*	0.03	0.97
Consumer Services	-0.09	2.65	-0.48	5.25	35.63*	7.68	49.48*	142.9*	0.04	0.88
Technology	-0.11	5.62	-0.24	4.70	84.55*	13.52**	118.03*	73.8*	0.08	0.81
Telecommunications	-0.01	3.63	-0.07	4.01	72.18*	15.03**	125.98*	24.7*	0.04	0.73
Utilities	0.07	2.25	-1.57	13.08	26.84*	4.09	36.58*	2635.3*	0.08	0.72

Notes: this table reports the basic statistics of return series, including mean (Mean), standard deviations (Std. dev.) skewness (Skew.), and kurtosis (Kurt.). ARCH test is the statistical test for conditional heteroscedasticity of order 6. LB and LB<sup>2</sup> are the Ljung-Box tests for autocorrelations of order 6 for the returns and for the squared returns. JB is the Jarque-Bera test for normality based on skewness and excess kurtosis. Corr. denotes the correlation coefficients. \*, \*\* and \*\*\* indicate the rejection of the null hypothesis of associated statistical tests at the 1%, 5% and 10% levels respectively.

Correlations between oil price changes and European sector returns are generally weak, and surprisingly they are all positive, except for the three following sectors: *Automobile and Parts*, *Food and Beverages*, and *Health Care*. This suggests that oil price increases over the last decade were likely to be seen as an indicator of higher expected economic growth and earnings. The sector *Oil and Gas* has the highest degree of comovement with oil (0.33), followed by the sector *Basic Materials* (0.12). Correlations between the European market index (DJ Stoxx 600) and sector returns are high on average. The *Personal and Household Goods* sector shows the highest correlation (0.97) and the sector *Food & Beverages* the lowest one (0.57).

### 3. Empirical analysis



We investigate the relationships between oil price changes and sector stock market returns in Europe over the last turbulent decade. We begin our analysis with the estimation of multifactor asset pricing models to investigate the sensitivities of the sector stock returns to oil price and European market changes, then we perform the Granger causality tests to examine their causal linkages, and finally we study cyclical comovements.

### **3.1 Sector returns, oil price changes and market sensitivities**

In this subsection, the analysis is conducted as follows. First, we estimate a conditional version of the European market model for each sector. Second, we examine a two-factor model by introducing oil unexpected returns into the market model. The objective is to investigate sector return sensitivities to oil price shocks. Finally, we test for asymmetric interactions between oil price changes and European sector returns.

#### *a) The market model*

A conditional version of the European market model can be written as follows (Model 1):<sup>2</sup>

$$r_{it} = a + c \times rdj_t + \varepsilon_{it} \quad (1)$$

$$\varepsilon_{it} \rightarrow f(0, h_{it})$$

$$h_{it}^2 = \alpha + \beta_k \times \sum_{k=1}^q \varepsilon_{i,t-1}^2 + \gamma_l \times \sum_{l=1}^p h_{i,t-1}^2$$

where

$r_{it}$  is the weekly stock returns for sector  $i$ ;

$rdj_t$  represents the European stock market returns;

$\varepsilon_{it}$  refers to a stochastic error term which is assumed to follow a GARCH( $q, p$ ) dynamics.

$p$  and  $q$  are explicitly determined according to commonly used information criteria.

$f(\cdot)$  is the density function of  $\varepsilon_{it}$ .

Model 1 is estimated for each of the considered sectors using the quasi-maximum likelihood (QML) method based on the Gaussian distribution. Here we also employ the Student's  $t$ -distribution to capture the distribution of sector returns because most series are highly skewed and exhibit significant excess kurtosis, leading to the rejection of normality.<sup>3</sup>

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<sup>2</sup> When estimating our market models, an AR(1) term is used wherever it appears to be significant.

<sup>3</sup> Note however that the use of the Student's  $t$ -distribution is motivated by comparative purpose and thus does not disprove the results from assuming the normal distribution since the QML estimator is consistent and asymptotically normal under certain regularity conditions, even if the normality assumption is violated (Bollerslev and Wooldridge, 1992).

We summarize the results in Table 3. As we can see, the estimates of model's parameters are somewhat similar, whatever the return distribution was used. However, according to the AIC information criteria, the model estimated using the Student's t-distribution shows superior results in 10 out of 12 cases. The evidence is mixed according to R-squared and LB criteria. The coefficients relating the sector returns to the European stock market returns (coefficients  $c$ ) are highly significant for all sectors. They vary from 0.50 (defensive sector) for *Food and Beverages* to 1.46 (offensive sector) for *Technology*. The R-squared coefficients range from 31% (*Food and Beverages*) to 82% (*Financials*). The models we estimated seem to satisfactorily fit the data. The ARCH and GARCH coefficients are significant. We further observe that in most cases, conditional volatility does not change very sharply as the ARCH coefficients are relatively small in size. By contrast, it tends to fluctuate gradually over time because of the large GARCH coefficients. Note also that the estimates coefficients  $\gamma$  and  $\beta$  satisfy the stationary conditions.

**Table 3. Estimation results of the European market model**

Sectors	Distribution	$a$	$c$	$\alpha$	$\beta_1$	$\gamma_1$	$\bar{R}^2$	AIC	ARCH test	$LB$	$LB^2$	$JB$
Automobiles and Parts	<i>Normal</i>	0.001 (0.001)	1.232* (0.041)	0.001* (0.000)	0.193* (0.057)	0.752* (0.061)	0.455	-4.558	4.900	6.219	5.314	83.115 <sup>+</sup>
	<i>Student-t</i>	0.000 (0.000)	1.238* (0.036)	0.001** (0.000)	0.184* (0.051)	0.755* (0.068)	0.455	-4.636	5.857	6.461	6.459	84.942 <sup>+</sup>
Financials	<i>Normal</i>	-0.001*** (0.001)	1.189* (0.025)	0.001* (0.000)	0.542* (0.120)	-	0.823	-6.184	2.357	7.154	2.322	26.839 <sup>+</sup>
	<i>Student-t</i>	-0.001** (0.000)	1.178* (0.013)	0.001* (0.000)	0.625* (0.223)	-	0.822	-6.205	2.211	6.766	2.198	27.020 <sup>+</sup>
Food and Beverages	<i>Normal</i>	0.001 (0.001)	0.504* (0.038)	0.000* (0.000)	0.214* (0.073)	-	0.313	-5.505	8.289	2.247	8.315	26.375 <sup>+</sup>
	<i>Student-t</i>	0.001*** (0.001)	0.511* (0.021)	0.000* (0.000)	0.295* (0.115)	-	0.311	-5.517	8.859	2.091	8.894	31.315 <sup>+</sup>
Oil and Gas	<i>Normal</i>	0.000 (0.001)	0.793* (0.047)	0.001* (0.000)	0.202* (0.090)	-	0.410	-4.652	3.626	6.732	3.552	12.877 <sup>+</sup>
	<i>Student-t</i>	0.000 (0.001)	0.799* (0.034)	0.001* (0.000)	0.206* (0.077)	-	0.410	-4.613	3.374	6.823	3.302	15.309 <sup>+</sup>
Health Care	<i>Normal</i>	0.000 (0.001)	0.596* (0.038)	0.001* (0.000)	0.066* (0.024)	0.927* (0.027)	0.363	-5.131	2.798	4.380	2.786	185.465 <sup>+</sup>
	<i>Student-t</i>	0.001 (0.001)	0.641* (0.026)	0.001* (0.000)	0.061* (0.026)	0.938* (0.019)	0.361	-5.201	3.066	4.826	3.030	264.367 <sup>+</sup>
Industrials	<i>Normal</i>	0.000 (0.001)	1.097* (0.025)	0.001* (0.000)	0.092* (0.024)	0.901* (0.022)	0.809	-6.107	1.769	2.901	1.800	25.343 <sup>+</sup>
	<i>Student-t</i>	0.000 (0.001)	1.078* (0.017)	0.001* (0.000)	0.090* (0.026)	0.903* (0.025)	0.802	-6.083	2.177	2.775	2.221	30.836 <sup>+</sup>
Basic Materials	<i>Normal</i>	0.002* (0.001)	1.100* (0.050)	0.000* (0.000)	0.522* (0.122)	-	0.520	-5.068	2.219	3.960	2.197	97.920 <sup>+</sup>
	<i>Student-t</i>	0.001** (0.001)	1.042* (0.029)	0.000* (0.000)	0.501* (0.119)	-	0.524	-5.094	1.992	3.864	1.976	141.832 <sup>+</sup>
Personal and Household Goods	<i>Normal</i>	0.001* (0.001)	0.896* (0.023)	0.001* (0.000)	0.081* (0.043)	0.893* (0.040)	0.750	-6.094	3.160	6.119	3.145	19.249 <sup>+</sup>
	<i>Student-t</i>	0.001* (0.001)	0.886* (0.016)	0.001* (0.000)	0.080* (0.022)	0.901* (0.028)	0.751	-6.105	2.928	6.033	3.683	22.386 <sup>+</sup>
Consumer Services	<i>Normal</i>	0.000 (0.001)	0.925* (0.018)	0.000* (0.000)	0.083* (0.035)	0.901* (0.036)	0.766	-6.092	2.243	5.596	2.323	76.473 <sup>+</sup>
	<i>Student-t</i>	0.000 (0.001)	0.917* (0.019)	0.000* (0.000)	0.062* (0.019)	0.920* (0.022)	0.766	-6.137	2.888	5.773	3.016	90.125 <sup>+</sup>
Technology	<i>Normal</i>	-0.000 (0.001)	1.460* (0.058)	0.001* (0.000)	0.080* (0.023)	0.901* (0.030)	0.661	-4.403	1.198	4.432	1.230	110.020 <sup>+</sup>

	<i>Student-t</i>	-0.000 (0.001)	1.395* (0.036)	0.001* (0.000)	0.080* (0.027)	0.900* (0.027)	0.656	-4.444	1.038	4.985	1.059	200.180 <sup>+</sup>
Telecommunica- tions	<i>Normal</i>	0.001 (0.001)	0.973* (0.036)	0.000* (0.000)	0.070* (0.014)	0.911* (0.013)	0.517	-4.761	2.661	13.350 <sup>++</sup>	2.665	72.107 <sup>+</sup>
	<i>Student-t</i>	-0.000 (0.001)	0.975* (0.035)	0.000* (0.000)	0.088* (0.025)	0.907* (0.021)	0.517	-4.801	3.192	12.645 <sup>++</sup>	3.150	102.450 <sup>+</sup>
Utilities	<i>Normal</i>	0.001** (0.001)	0.603* (0.029)	0.001** (0.001)	0.165* (0.063)	0.721* (0.079)	0.528	-5.554	5.710	6.039	5.961	101.720 <sup>+</sup>
	<i>Student-t</i>	0.001** (0.001)	0.604* (0.023)	0.001** (0.001)	0.114* (0.051)	0.757* (0.111)	0.527	-5.999	7.132	5.949	7.321	142.970 <sup>+</sup>

Notes: this table reports the results from estimating the European market model for sector returns. Numbers in parenthesis are robust standard errors.  $LB$  and  $LB^2$  are the Ljung-Box tests for autocorrelation of order 6 for the standardised residuals and for the squared residuals. ARCH test is the LM ARCH test for conditional heteroscedasticity of order 6.  $JB$  is the Jarque-Bera test for normality. AIC is the Akaike Information Criterion. For Food and Beverages, Basic Materials, and Industrials sectors, the model is estimated with an AR(1) because the latter is significant. We also tested for GARCH effect in the mean equation, but the associated coefficients are not significant. The orders for the GARCH model are determined based on information criteria. The degree of freedom  $\nu$  for the Student's t-distribution is significantly higher than 2 in all cases, suggesting that the distribution of the standardized errors departs significantly from normality.

\*, \*\* and \*\*\* indicate the significance of coefficients at the 1%, 5% and 10% levels respectively.

<sup>+</sup>, <sup>++</sup> and <sup>+++</sup> indicate the rejection of the null hypothesis of statistical tests at the 1%, 5% and 10% levels respectively.

The Jarque-Bera statistics in Table 3 are considerably lower than those for the return series (Table 2). For instance, the JB statistic decreases from 5930.1 (*Automobile and Parts*), 2635.3 (*Utilities*) and 913.1 (*Oil and Gas*) to 83.1, 101.7 and 12.9 respectively. However, the normality hypothesis is still rejected indicating that the unconditional distribution of the conditional GARCH process is not sufficiently fat-tailed to accommodate the excess kurtosis in the data. This result justifies the use of the QML estimation method and the Student's t-distribution.

Finally, we also test for ARCH effects as well as for the absence of autocorrelation in the standardised residuals and in the squared residuals. The results indicate no serial correlations and heteroscedastic effects in the residuals, thus leading us to conclude that the model specification we use is flexible enough to capture the dynamics of returns.

*b) The two-factor "market and oil" model<sup>4</sup>*

Let us now consider an augmented version of the previous European market model by introducing the unexpected change in oil prices into Equation (1). This specification permits to assess the sensitivities of sector returns to oil price shocks and has the following form (Model 2):<sup>5</sup>

$$r_{it} = a + b \times roil_t^u + c \times rdj_t + \varepsilon_{it} \quad (2)$$

$$\varepsilon_{it} \rightarrow f(0, h_{it})$$

$$h_{it}^2 = \alpha + \beta_k \times \sum_{k=1}^q \varepsilon_{i,t-1}^2 + \gamma_l \times \sum_{l=1}^p h_{i,t-1}^2$$

where  $roil_t^u$  is the unexpected change in oil prices, measured as the difference between the observed oil price change and the expected value of oil price change using the following regression model  $roil_{it} = \mu + \sum_{l=1}^k \lambda_l roil_{i,t-l} + \zeta_{it}$

Obviously, the definition of unexpected changes in oil prices we retain in this paper would mean that the impact of previous oil price changes on stock returns is implicitly included in Equation (2). In this regard, the estimation results can be seen as a sort of causality tests

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<sup>4</sup> We have also tested other multifactor models in which the relationships between oil and stock prices are controlled for by using other potential risk factors of stock returns. These factors include the changes in short-term interest rates, the changes in consumer price index, and the changes in industrial production. Since the obtained results are very similar to our basic two-factor model as described by Equation (2), they are not reported here for concision purpose, but entirely available under request.

<sup>5</sup> The suitability of two-factor "market and oil" pricing models, similar to the one we use in this paper, was empirically investigated in several past papers (see, e.g., Faff and Brailsford, 2000 and Arouri and Fouquau, 2009).

between oil and stock returns. Based on information criteria, we retain three lags ( $k = 3$ ) in order to appropriately remove autocorrelations in oil returns.

**Table 4. Estimation results of the two-factor “market and oil” model**

Sectors	Distribution	$a$	$b$	$c$	$\bar{R}^2$	AIC	ARCH test	$LB$	$LB^2$	$JB$
Automobile and Parts	<i>Normal</i>	0.001 (0.001)	-0.031 (0.029)	1.232* (0.042)	0.459	-4.648	5.948	6.027	6.482	80.219 <sup>+</sup>
	<i>Student-t</i>	0.000 (0.001)	-0.033*** (0.020)	1.238* (0.036)	0.458	-4.496	7.031	6.254	7.702	82.615 <sup>+</sup>
Financials	<i>Normal</i>	0.000 (0.000)	0.029** (0.015)	1.170* (0.017)	0.851	-6.341	2.823	6.804	2.666	32.726 <sup>+</sup>
	<i>Student-t</i>	-0.000 (0.000)	0.031** (0.017)	1.165* (0.015)	0.851	-6.350	2.211	6.199	2.335	49.002 <sup>+</sup>
Food and Beverages	<i>Normal</i>	0.001 (0.001)	-0.049* (0.014)	0.553* (0.026)	0.336	-5.523	8.289	1.538	7.912	19.522 <sup>+</sup>
	<i>Student-t</i>	0.001 (0.001)	-0.049* (0.012)	0.556* (0.022)	0.337	-5.534	8.859	1.334	6.978	22.859 <sup>+</sup>
Oil and Gas	<i>Normal</i>	-0.000 (0.001)	0.206* (0.019)	0.820* (0.035)	0.505	-4.851	4.830	7.376	4.797	21.222 <sup>+</sup>
	<i>Student-t</i>	0.000 (0.001)	0.204* (0.019)	0.818* (0.033)	0.505	-4.864	4.843	7.818	4.776	15.309 <sup>+</sup>
Health Care	<i>Normal</i>	-0.001 (0.001)	-0.075* (0.015)	0.637* (0.026)	0.404	-5.178	2.469	8.265	2.459	151.376 <sup>+</sup>
	<i>Student-t</i>	0.000 (0.000)	-0.083* (0.018)	0.594* (0.037)	0.405	-5.243	2.567	7.259	2.548	215.882 <sup>+</sup>
Industrials	<i>Normal</i>	0.001 (0.001)	0.016*** (0.009)	1.097* (0.025)	0.831	-6.115	1.899	2.788	1.800	25.687 <sup>+</sup>
	<i>Student-t</i>	0.000 (0.001)	0.0015** (0.008)	1.079* (0.017)	0.832	-6.128	2.436	2.650	2.340	33.006 <sup>+</sup>
Basic Materials	<i>Normal</i>	0.002** (0.001)	0.032** (0.015)	1.085* (0.038)	0.600	-5.069	2.113	3.553	2.137	78.634 <sup>+</sup>
	<i>Student-t</i>	0.002** (0.001)	0.025*** (0.014)	1.050* (0.028)	0.601	-5.095	1.891	3.252	1.893	115.556 <sup>+</sup>
Personal and Household Goods	<i>Normal</i>	0.001*** (0.001)	-0.035 (0.023)	0.899* (0.023)	0.762	-6.110	7.351	8.040	6.680	13.710 <sup>+</sup>
	<i>Student-t</i>	0.001*** (0.001)	-0.035 (0.024)	0.891* (0.016)	0.763	-6.117	6.798	7.867	6.148	15.224 <sup>+</sup>
Consumer Services	<i>Normal</i>	0.000 (0.000)	0.023* (0.012)	0.927* (0.018)	0.788	-6.093	2.181	5.076	2.289	92.179 <sup>+</sup>
	<i>Student-t</i>	0.000 (0.000)	0.025** (0.013)	0.919* (0.018)	0.788	-6.143	3.110	5.367	3.276	107.801 <sup>+</sup>
Technology	<i>Normal</i>	0.001 (0.001)	-0.043** (0.022)	1.452* (0.059)	0.690	-4.486	1.023	4.189	1.032	120.890 <sup>+</sup>
	<i>Student-t</i>	-0.000 (0.000)	-0.053** (0.021)	1.387* (0.036)	0.694	-4.511	0.926	4.713	0.657	231.910 <sup>+</sup>
Telecommunications	<i>Normal</i>	-0.001 (0.001)	-0.026 (0.021)	0.975* (0.032)	0.529	-4.781	1.904	11.751+++	1.965	86.954 <sup>+</sup>
	<i>Student-t</i>	-0.000 (0.001)	-0.018 (0.022)	0.966* (0.034)	0.528	-4.810	2.416	9.806+++	2.423	149.860 <sup>+</sup>
Utilities	<i>Normal</i>	0.001 (0.001)	0.008 (0.013)	0.603* (0.029)	0.529	-5.558	5.718	5.246	6.086	98.310 <sup>+</sup>
	<i>Student-t</i>	0.001** (0.001)	0.004 (0.012)	0.605* (0.022)	0.529	-5.602	7.187	5.470	7.101	141.85 <sup>+</sup>

Notes: this table reports the results from estimating the tow-factor “market and oil” model for sector returns. Numbers in parenthesis are robust standard errors.  $LB$  and  $LB^2$  are the Ljung-Box tests for autocorrelation of order 6 for the standardised residuals and for the squared residuals. ARCH test is the LM ARCH test for conditional heteroscedasticity of order 6.  $JB$  is the Jarque-Bera test for normality. AIC is the Akaike Information Criterion. For Food and Beverages, Health Care, Basic Materials, and Industrials sectors, the model is estimated with an AR(1) because the latter is significant. We also tested for GARCH effect in the mean equation, but the associated coefficients are not significant. The orders for the GARCH model are determined based on information criteria. The degree of freedom  $\nu$  for the Student’s t-distribution is significantly higher than 2 in all cases, suggesting that the distribution of the standardized errors departs significantly from normality. The GARCH coefficients are not reported here in order to preserve space, but they are similar to those reported in Table 3.

\*, \*\* and \*\*\* indicate the significance of coefficients at the 1%, 5% and 10% levels respectively.

<sup>+</sup>, <sup>++</sup> and <sup>+++</sup> indicate the rejection of the null hypothesis of statistical tests at the 1%, 5% and 10% levels respectively.

The estimation results are summarized in Table 4.<sup>6</sup> The coefficients relating the return series to oil price changes (coefficients *b*) are significant in eight cases, indicating significant short-term effects of oil price fluctuations on European sector stock returns. Oil price increases negatively affect sector returns in three cases (*Food and Beverages*, *Health Care* and *Technology*), and positively in five cases (*Financials*, *Oil and Gas*, *Industrials*, *Basic Materials*, and *Consumer Services*), which is consistent with the correlations reported in Table 2. This confirms our intuition that, while higher oil prices imply lower stock returns for some industries due to higher production and transportation costs and lower corporate earnings, increases in oil prices over the last decade also reflect the increases in world demand for oil in response to periods of high economic growth, and thus lead to positive stock returns for other sectors. Additionally, the sign of the oil-stock price relationships is also likely to be dependent on the capacity of the industry to transfer oil price shocks to other economic entities, through for example hedging contracts on commodity derivatives markets, and thus to minimise the impact of these shocks on its profitability. Finally, our results show that there is no relationship between oil price changes and stock returns for three European sectors (*Personal and Household Goods*, *Telecommunications*, and *Utilities*), whereas for the *Automobile and Parts* industry a negative weak link is obtained.

It is equally important to note that whenever oil price changes are significant, the two-factor “market and oil” model outperforms the market model as the AIC and R-squared scores in Table 4 are respectively smaller and larger than those in Table 3.

Summarizing all, our analysis shows strong linkages between oil price changes and most European sector returns over the period under consideration. The sign and intensity of these linkages differ from one sector to another. In the following sub-section, we test for asymmetries in the responses of European sector returns to oil price shocks.

#### *c) Asymmetric reaction to oil shocks*

Some recent papers have shown that the link between oil and economic activity is not entirely linear and that negative oil price shocks (price increases) tend to have larger impacts on growth than positive shocks do (Hamilton, 2003; Lardic and Mignon, 2006; Zhang, 2008; Cologni and Manera, 2009). One should expect that oil price changes equally affect stock markets in an asymmetric fashion. To empirically test for asymmetry in the reaction of

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<sup>6</sup> We do not report the estimates of the GARCH coefficients as they are very similar to those reported in Table 3. Note also that when estimating the two-factor “market and oil” model, an autoregressive term AR(1) is used wherever it shows up as being significant.

European sector returns to oil price shocks, we rely on the estimation of the following asymmetric multifactor model (Model 3):

$$r_{it} = a + b^+ D_t \times roil_t^u + b^- (1 - D_t) \times roil_t^u + c \times rdj_t + \varepsilon_{it} \quad (3)$$

$$\varepsilon_{it} \rightarrow f(0, h_{it})$$

$$h_{it}^2 = \alpha + \beta_k \times \sum_{k=1}^q \varepsilon_{i,t-1}^2 + \gamma_l \times \sum_{l=1}^p h_{i,t-1}^2$$

where  $D_t$  is a dummy variable taking a value of one if unexpected change in oil price is positive and zero if it is negative.<sup>7</sup> Accordingly,  $b^+$  and  $b^-$  are the coefficients corresponding to increases and decreases in unexpected oil price respectively. There is no asymmetry if  $b^+$  and  $b^-$  are not statistically different from each other, which requires us to test the null hypothesis of coefficient equality,  $b^+ = b^-$ . We are also interested in testing for the null hypothesis of non asymmetry and nonsensitivities to oil price increases and decreases,  $b^+ = b^- = 0$ .

Our main empirical results are summarized in Table 5. Wald tests show that the hypothesis  $b^+ = b^- = 0$  is rejected mostly at the 1% level in nine cases, which confirms the significance of oil price shocks as a factor affecting sector returns in Europe. Oil price changes do not significantly affect stock returns in the *Automobile* and *Parts*, *Telecommunications*, and *Utilities* sectors. This is in line with our findings reported in Table 4. The only exception is for *Personal* and *Household Goods* for which no significant reaction to oil shocks is observed when the symmetric asset pricing model was used. Indeed, asymmetric results in Table 5 show that stock returns in this sector react negatively to unexpected oil price increases and negatively to expected oil price decreases. Wald tests confirm this finding and show that the hypothesis  $b^+ = b^-$  is rejected for *Personal* and *Household Goods* as well as for two other sectors (*Food* and *Beverages*, and *Health Care*). This hypothesis is also weakly rejected at 10% for *Basic Materials*. For all these industries, reactions to oil price changes are asymmetric.

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<sup>7</sup> Note that we have also estimated the asymmetric multifactor model with an EGARCH(1,1) process, but information criteria lead us to prefer the GARCH( $q,p$ ) specification we report in this paper.



**Table 5. Estimation results of the asymmetric asset pricing model**

Sectors	Distribution	$b^+$	$b^-$	$b^+ = b^- = 0$	$b^+ = b^-$	$\bar{R}^2$	AIC	ARCH Test	LB	LB <sup>2</sup>	JB
Automobile and Parts	<i>Normal</i>	-0.026 (0.040)	-0.047 (0.039)	3.041 [0.218]	0.097 [0.755]	0.460	-4.645	6.100	6.040	6.686	79.930 <sup>+</sup>
	<i>Student-t</i>	-0.018 (0.038)	-0.039 (0.024)	2.026 [0.363]	0.099 [0.752]	0.459	-4.493	7.199	6.281	7.921	82.331 <sup>+</sup>
Financials	<i>Normal</i>	0.016 (0.018)	0.040** (0.016)	11.475 [0.003]	0.646 [0.422]	0.851	-6.338	2.962	6.886	2.808	31.164 <sup>+</sup>
	<i>Student-t</i>	0.019 (0.017)	0.042* (0.017)	14.650 [0.001]	0.639 [0.423]	0.851	-6.357	2.211	6.302	2.474	46.543 <sup>+</sup>
Food and Beverages	<i>Normal</i>	-0.071* (0.023)	0.031 (0.020)	18.354 [0.000]	12.299 [0.000]	0.343	-5.532	7.109	1.522	6.545	16.032 <sup>+</sup>
	<i>Student-t</i>	-0.061** (0.025)	0.039*** (0.021)	15.681 [0.000]	13.339 [0.000]	0.343	-5.541	6.768	1.336	6.355	22.859 <sup>+</sup>
Oil and Gas	<i>Normal</i>	0.190* (0.042)	0.221* (0.029)	133.712 [0.000]	0.239 [0.624]	0.505	-4.828	4.419	7.122	4.660	21.300 <sup>+</sup>
	<i>Student-t</i>	0.190* (0.041)	0.218* (0.030)	123.499 [0.000]	0.233 [0.629]	0.505	-4.841	4.732	7.590	4.661	25.141 <sup>+</sup>
Health Care	<i>Normal</i>	-0.022 (0.032)	-0.134* (0.027)	36.831 [0.000]	4.579 [0.032]	0.457	-5.192	2.216	6.172	2.874	150.063 <sup>+</sup>
	<i>Student-t</i>	-0.002 (0.028)	-0.135* (0.026)	33.167 [0.000]	8.250 [0.004]	0.457	-5.291	2.785	6.292	2.750	206.987 <sup>+</sup>
Industrials	<i>Normal</i>	-0.014 (0.009)	0.019*** (0.012)	11.130 [0.003]	0.974 [0.327]	0.831	-6.104	1.741	2.837	1.646	26.679 <sup>+</sup>
	<i>Student-t</i>	-0.023 (0.016)	0.018*** (0.011)	11.606 [0.002]	1.601 [0.206]	0.833	-6.128	2.355	2.703	2.244	34.203 <sup>+</sup>
Basic Materials	<i>Normal</i>	-0.015 (0.032)	0.073** (0.029)	16.282 [0.000]	2.806 [0.093]	0.674	-5.071	2.118	3.490	2.224	75.543 <sup>+</sup>
	<i>Student-t</i>	-0.017 (0.029)	0.062** (0.028)	19.058 [0.000]	2.709 [0.099]	0.674	-5.123	1.930	3.208	1.929	98.098 <sup>+</sup>
Personal and Household Goods	<i>Normal</i>	-0.077* (0.021)	0.051** (0.026)	16.580 [0.000]	5.026 [0.025]	0.793	-6.136	7.132	8.019	6.523	11.555 <sup>+</sup>
	<i>Student-t</i>	-0.074* (0.019)	0.051** (0.025)	15.133 [0.000]	6.030 [0.014]	0.793	-6.222	7.011	7.768	6.099	13.054 <sup>+</sup>
Consumer Services	<i>Normal</i>	0.032*** (0.019)	0.016 (0.016)	6.853 [0.032]	0.265 [0.696]	0.788	-6.090	2.174	5.083	2.284	90.565 <sup>+</sup>
	<i>Student-t</i>	0.031 (0.021)	0.018 (0.019)	5.968 [0.051]	0.185 [0.666]	0.788	-6.140	2.907	5.123	2.934	103.886 <sup>+</sup>
Technology	<i>Normal</i>	-0.010 (0.032)	-0.072** (0.035)	5.660 [0.059]	1.160 [0.281]	0.691	-4.487	1.184	4.121	1.200	113.765 <sup>+</sup>
	<i>Student-t</i>	-0.016 (0.034)	-0.088* (0.031)	11.722 [0.003]	1.620 [0.203]	0.694	-4.518	0.909	4.332	0.765	208.123 <sup>+</sup>
Telecommunications	<i>Normal</i>	0.031 (0.042)	-0.062 (0.047)	1.766 [0.413]	1.361 [0.243]	0.530	-4.386	1.768	11.618 <sup>+++</sup>	1.858	81.098 <sup>+</sup>
	<i>Student-t</i>	-0.001 (0.042)	-0.035 (0.040)	0.947 [0.622]	0.244 [0.621]	0.529	-4.427	1.978	10.112 <sup>+++</sup>	2.098	145.768 <sup>+</sup>
Utilities	<i>Normal</i>	0.001 (0.001)	-0.008 (0.013)	0.678 [0.712]	0.495 [0.481]	0.530	-5.556	5.871	5.192	6.332	88.786 <sup>+</sup>
	<i>Student-t</i>	-0.005 (0.024)	-0.0041 (0.022)	0.122 [0.940]	0.001 [0.977]	0.529	-5.599	7.875	5.445	6.654	132.988 <sup>+</sup>

Notes: this table reports the results from estimating the asset pricing model with asymmetric reaction of sector returns to oil price shocks. Numbers in parenthesis are robust standard errors. In columns 5 and 6, we report empirical statistics of the Wald tests and their associated  $p$ -values in brackets.  $LB$  and  $LB^2$  are the Ljung-Box tests for autocorrelation of order 6 for the standardised residuals and for the squared residuals. ARCH test is the LM ARCH test for conditional heteroscedasticity of order 6.  $JB$  is the Jarque-Bera test for normality. AIC is the Akaike Information Criterion. For Food and Beverages, Health Care, Basic Materials, and Industrials sectors, the model is estimated with an AR(1) because the latter is significant. We also tested for GARCH effect in the mean equation, but the associated coefficients are not significant. The orders for the GARCH model are determined based on information criteria. The degree of freedom  $\nu$  for the Student's  $t$ -distribution is significantly higher than 2 in all cases, suggesting that the distribution of the standardized errors departs significantly from normality. The GARCH coefficients are not reported for concision purpose, but they are similar to those reported in Table 3. \*, \*\* and \*\*\* indicate the significance of coefficients at the 1%, 5% and 10% levels respectively. +, ++ and +++ indicate the rejection of the null hypothesis of statistical tests at the 1%, 5% and 10% levels respectively.

Finally, it is worth noting that according the different criteria used to chose the most appropriate model (R-squared, AIC, LB, etc.), the asymmetric asset pricing model appears to be the best one when there is presence of asymmetry in the relationship between oil and stock returns, *i.e.*, for these four industries *Personal* and *Household Goods*, *Food and Beverages*, *Basic Materials*, and *Health Care*. Also, the models incorporating oil returns (*i.e.*, Models 2 3) shows superior results to the market model (*i.e.*, Model 1) for all the industries for which oil price changes are priced, the only exceptions being *Telecommunications*, and *Utilities* sectors. All in all, these findings suggest that oil price changes play a significant role in explaining stock returns in most European industries and that there are evidence to show that some sector returns respond asymmetrically to the impact of oil price changes.

### 3.2 Causality tests

In order to further examine the relationships between oil price changes and sector stock returns in Europe, we proceed with testing for Granger causality between return series. Results are reported in Table 6. Since some variables as well as their bilateral effects are very sensitive to the selected number of lags in the analysis, this test is implemented for different lags.

The results show that there is bidirectional causality between oil price changes and DJ Stoxx returns. Indeed, DJ Stoxx returns granger-cause changes in oil prices at the 10% level for from one to three autoregressive lags, whereas oil price shocks granger-cause changes in DJ Stoxx returns at the 5% level for all lags, except the first one. Similar results are obtained for different lags with regard to *Automobile and Parts*, *Food and Beverages*, *Oil and Gas*, *Industrials*, *Personal* and *Household Goods*, *Consumer Services* and *Utilities* industries. Unidirectional Granger causality from oil to stock returns is significant for *Financials* (at the 10% level for lags 2 and 8), *Health Care* (at the 10% level only for lag 2), and *Technology* (at the 10% level for lag 2 and at the 5% for lags 8, 10 and 12), while causality from stock returns to oil is only found to be significant for *Basic Materials* (at the 1% for all lags). Finally, there is absence of significant causality between oil price changes and stock returns in *Telecommunications* sector.

Taken together, the results of our causality tests corroborate our previous findings and suggest significant interactions between oil prices and stock prices, except for *Telecommunications* stocks. These results are interesting at least for two reasons. First, they imply some predictability in oil and European stock price dynamics. Second, in contrast to several works in the extant literature which assume the exogeneity of oil prices with respect to macroeconomic and financial variables, we document that a reverse relationship may exist,

*i.e.*, changes in some European sector returns do significantly affect world oil prices. These findings are consistent with the results established by some recent papers (Ewing and Thompson, 2007; Kilian and Vega, 2008; Lescaroux and Mignon, 2008).

**Table 6. Results of the Granger causality tests**

Lags	1	2	3	4	6	8	10	12
DJ Stoxx Europe								
S→O	<b>0.073</b>	<b>0.075</b>	0.111	<b>0.099</b>	0.216	0.112	0.104	0.362
O→S	0.253	<b>0.035</b>	<b>0.055</b>	<b>0.097</b>	0.159	0.172	0.185	0.372
Automobile and Parts								
S→O	<b>0.016</b>	<b>0.019</b>	<b>0.037</b>	<b>0.062</b>	<b>0.049</b>	<b>0.041</b>	<b>0.035</b>	0.138
O→S	0.133	<b>0.007</b>	<b>0.007</b>	<b>0.006</b>	<b>0.005</b>	<b>0.011</b>	<b>0.023</b>	<b>0.048</b>
Financials								
S→O	0.345	0.137	0.216	0.191	0.363	0.234	0.140	0.436
O→S	0.137	<b>0.076</b>	0.117	0.155	0.277	<b>0.077</b>	0.150	0.272
Food and Beverages								
S→O	<b>0.055</b>	<b>0.023</b>	<b>0.005</b>	<b>0.002</b>	<b>0.004</b>	<b>0.003</b>	<b>0.004</b>	<b>0.015</b>
O→S	0.102	<b>0.064</b>	0.150	<b>0.086</b>	0.119	0.142	0.178	0.315
Oil and Gas								
S→O	<b>0.068</b>	<b>0.060</b>	<b>0.058</b>	<b>0.069</b>	<b>0.056</b>	<b>0.082</b>	<b>0.063</b>	0.160
O→S	0.635	<b>0.040</b>	0.136	0.144	0.306	0.341	0.470	0.487
Health Care								
S→O	0.632	0.441	0.250	0.324	0.456	0.158	0.161	0.468
O→S	0.128	<b>0.055</b>	0.149	0.272	0.510	0.499	0.377	0.445
Industrials								
S→O	<b>0.030</b>	<b>0.022</b>	<b>0.043</b>	<b>0.025</b>	<b>0.060</b>	<b>0.032</b>	<b>0.039</b>	0.180
O→S	0.807	<b>0.057</b>	<b>0.063</b>	0.106	0.154	<b>0.087</b>	<b>0.072</b>	0.166
Basic Materials								
S→O	<b>0.005</b>	<b>0.004</b>	<b>0.009</b>	<b>0.007</b>	<b>0.003</b>	<b>0.001</b>	<b>0.001</b>	<b>0.008</b>
O→S	0.625	0.264	0.426	0.195	0.338	0.251	0.280	0.404
Personal and Household Goods								
S→O	<b>0.070</b>	<b>0.042</b>	<b>0.074</b>	<b>0.085</b>	0.163	<b>0.078</b>	<b>0.060</b>	0.246
O→S	<b>0.069</b>	<b>0.097</b>	0.144	0.259	0.343	0.566	0.308	0.451
Consumer Services								
S→O	<b>0.056</b>	<b>0.038</b>	<b>0.063</b>	0.116	0.308	0.155	0.127	0.303
O→S	0.131	<b>0.069</b>	<b>0.090</b>	0.154	0.184	0.247	0.141	0.285
Technology								
S→O	0.529	0.803	0.903	0.400	0.454	0.469	0.532	0.852
O→S	0.324	<b>0.084</b>	0.112	0.208	0.112	<b>0.016</b>	<b>0.005</b>	<b>0.020</b>
Telecommunications								
S→O	0.582	0.811	0.902	0.606	0.791	0.807	0.771	0.909
O→S	0.607	0.116	0.222	0.233	0.196	0.213	0.149	0.247
Utilities								
S→O	<b>0.046</b>	<b>0.091</b>	0.138	0.240	0.393	0.387	0.313	0.531
O→S	0.673	<b>0.087</b>	0.212	0.322	0.597	0.768	0.871	0.943

Notes: The Granger tests are based on a linear VAR( $p$ ) model, where  $p$  is equal to 1, 2, 3, 4, 6, 8, 10 and 12, respectively. The table provides the  $p$ -values of rejection of the null hypothesis.

S→O is the null hypothesis of no Granger causality from stock market returns to oil price changes.

O→S is the null hypothesis of no Granger causality from oil price changes to stock market returns.

### **3.3 Cyclical correlations between oil and stock markets**

To the extent that variations in macroeconomic fundamentals may influence the direct short-run linkages between oil price changes and sector stock returns in Europe, it is relevant to investigate the cyclical correlations between variables of interest. To do so, we follow the methodology introduced by Serletis and Shahmoradi (2005) and applied by several papers to study the links between energy prices and economics activity (Ewing and Thompson, 2007; Lescaroux and Mignon, 2008). First, the Hodrick-Prescott (HP) filter is employed to

decompose each time-series variable in our study into long-run and business cycle components. We then compute the cross-correlations between the cyclical component of oil price ( $coil_t$ ) and that of sector stock market indices ( $cstock_t$ ). We denote these correlations by  $\rho(j)$  and they are computed for  $j = 0, \pm 1, \pm 2, \pm 3, \pm 4, \pm 5$  and  $\pm 6$ . Therefore, the cyclical correlations permit to assess the linkages that may exist between oil price and stock markets over the business cycle. They enable, in particular, the investigation of the dynamics of the short-run component comovements by providing information about both their strength and their synchronization. Following Serletis and Shahmoradi (2005), and Ewing and Thompson (2007), we consider that the two cyclical components are strongly correlated, weakly correlated, or uncorrelated for a shift  $j$  based on  $0.23 \leq |\rho(j)| < 1$ ,  $0.10 \leq |\rho(j)| < 0.23$ ,  $0 \leq |\rho(j)| < 0.10$ , respectively. If  $|\rho(j)|$  is high for a positive, zero, or negative value of  $j$ , then the cycle of oil prices is leading the cycle of stock markets by  $j$  periods, is synchronous, or is lagging behind the cycle of stock markets by  $j$  periods, respectively.

**Table 7. Cyclical correlations of oil prices with stock market indices**

Period $j$	-6	-5	-4	-3	-2	-1	0	1	2	3	4	5	6
DJ Stoxx	0.024	0.051	0.087	0.107	0.119	0.118	0.116	0.070	0.023	-0.011	-0.042	-0.080	-0.907
Europe													
Automobile and Parts	0.020	0.026	0.046	0.008	0.023	-0.064	-0.000	-0.170	-0.200	-0.204	<b>-0.243</b>	<b>-0.257</b>	<b>-0.282</b>
Financials	-0.073	-0.066	-0.044	-0.035	-0.025	-0.031	-0.187	-0.202	-0.206	-0.196	-0.183	-0.181	-0.169
Food and Beverages	0.023	-0.012	-0.045	-0.074	-0.093	-0.127	-0.158	-0.166	-0.164	-0.169	<b>-0.256</b>	-0.161	-0.167
Oil and Gas	<b>0.281</b>	<b>0.301</b>	<b>0.330</b>	<b>0.356</b>	<b>0.382</b>	<b>0.352</b>	<b>0.385</b>	<b>0.320</b>	<b>0.271</b>	<b>0.231</b>	0.196	0.151	0.105
Health Care	-0.193	-0.202	-0.203	<b>-0.294</b>	-0.191	-0.194	-0.183	-0.153	-0.128	-0.101	-0.073	-0.055	-0.036
Industrials	0.096	0.088	0.089	0.079	0.069	0.043	0.016	-0.014	-0.041	-0.054	-0.069	-0.090	-0.108
Basic Materials	<b>0.422</b>	<b>0.430</b>	<b>0.455</b>	<b>0.465</b>	<b>0.452</b>	<b>0.423</b>	<b>0.422</b>	<b>0.360</b>	<b>0.304</b>	<b>0.263</b>	0.222	0.180	0.131
Personal and Household Goods	0.032	0.014	-0.001	-0.025	-0.048	-0.086	-0.120	-0.155	-0.151	-0.157	-0.162	-0.173	-0.178
Consumer Services	-0.086	-0.097	-0.100	-0.108	-0.118	-0.138	-0.164	-0.182	-0.195	-0.200	-0.207	-0.216	-0.217
Technology	0.038	0.034	0.037	0.035	0.036	0.036	0.033	0.026	0.018	0.012	0.005	0.000	0.005
Telecom-munications	-0.020	-0.021	-0.024	-0.030	-0.031	-0.031	-0.034	-0.039	-0.046	-0.050	-0.052	-0.052	-0.049
Utilities	0.002	0.001	0.011	0.016	0.024	0.025	0.011	0.002	-0.024	-0.033	-0.040	-0.065	-0.095

Notes: this table provides the cyclical correlations between oil price and stock market prices measured by  $\rho(j) = \rho(coil_t, cstock_{t+j})$ . Bold type indicates high absolute value correlations.

The results for leads and lags from 1 to 6 are shown in Table 7. They globally confirm previous results. Positive weak-cyclical correlations are observed for the DJ Stoxx 600 Index

for negative values of  $j$ , suggesting that oil prices are pro-cyclical and lagging the European stock market generally after few weeks over the last decade. As we expect, oil prices and stock markets are strongly and contemporaneously correlated for *Oil and Gas*, and *Basic Materials* industries. Furthermore, for two sectors (*Automobile and Parts*, and *Food and Beverages*) highly negative cyclical correlations are observed for positive values of  $j$ , indicating that oil prices are counter-cyclical and leading the stock markets generally by few weeks ahead in these sectors. Weak correlations are obtained for four sectors: *Health Care*, *Industrials*, *Personal Goods* and *Consumer Services*. Surprisingly, we find no significant cyclical correlations between oil prices and stock market returns in *Technology* and *Utilities* industries, but as expected no correlation is also found for *Telecommunications* stocks.

Overall, our analysis shows significant relationships between most sector returns in Europe and oil price changes. However, there is some asymmetry in the reaction of some sectors to oil price shocks, and in all cases the value and sign of the sensitivities of stock returns to oil price changes vary significantly across sectors. In what follows, we carry out some robustness checks.

### **3.4 Robustness checks**

To check for the robustness of our empirical results, the following changes are made. First, given previous evidence in the literature on the significance of world market risk and exchange rate risk within international asset pricing model settings, especially in developed stock markets such as European markets, we re-estimate our empirical models in the presence of the world market returns and the euro-dollar exchange rate changes. Second, we re-estimate our models using different sample periods and data frequencies. Finally, our empirical results are re-examined using real oil and stock prices instead of nominal prices in order to take into account the effects of inflation.<sup>8</sup>

#### *a) Effects of exchange rates*

Oil is commonly priced in US dollars in world oil markets, and thus the latter play an important role in the energy industry to the extent that the dollar exchange rate affects directly the price at which producers and consumers perceive oil and oil-related products. Therefore, it is interesting to consider the potential links between the dollar exchange rates and the oil prices in our empirical investigation. In the literature, the link between the dollar and oil prices has been examined at theoretical levels by, among others, Krugman (1980) and Rogoff (1991), and at empirical levels by, among others, Zhou (1995), and Dibooglu (1995). The

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<sup>8</sup> In order to preserve space, the results of robustness tests are not reported here, but available upon request.

findings of these studies suggest the existence of a positive relationship between the two variables, *i.e.*, a rise in oil prices is linked to a dollar appreciation.

Moreover, recent papers show that exchange rate risk is internationally priced (De Santis *et al.*, 2003 and references therein). In this schema of things, the statistical significance of oil price changes in Models 2 and 3 could be due to the failure to include the euro-dollar exchange rate changes. To shed light on this issue, we re-estimate Models 2 and 3 in presence of the euro-dollar exchange rate changes as a risk factor. The results indicate that changes in exchange rate have significant effects on most European sector returns. Meanwhile, the relationships between oil price shocks and European sector returns remain mainly unchanged: oil price changes positively affect sector returns in five cases (*Financials*, *Oil and Gas*, *Industrials*, *Basic Materials*, and *Consumer Services*) and negatively in four cases (*Automobile and Parts*, *Food and Beverages*, *Health Care* and *Technology*). Strong signs of asymmetric reaction to oil price changes have been found, in particular for *Food and Beverages*, and *Personal and Household Goods* stock returns.

#### *b) Effects of world stock market*

If European markets are integrated into world stock market, it is reasonable to believe that European stock returns are exposed to the world market risk. This leads us to test for the significance of oil price changes within an asset pricing model that allows for world market risk to affect stock returns as a robustness check. The inclusion of this global factor is motivated by the fact that if oil price changes and world market returns are correlated, tests based on models as in Equations (2) and (3) may result in a spurious significance of the oil price changes because of the failure to account for global factors.

More precisely, we re-estimate augmented versions of Models 2 and 3 which incorporate the part of MSCI world market returns, unexplained by the DJ Stoxx index returns and Brent oil price changes. We find that the global factor, independent from European and oil factors are significant for several European sectors. However, the significance of the oil price changes seems to be unaffected after the introduction of the global factor.

#### *c) Alternative sample periods and frequencies*

The DJ Stoxx sector indices are launched in 1998, but their daily, weekly and monthly frequencies were built back to December 31, 1986. As an exercise of the robustness check, we test the sensitivity of our results to the choice of the sample periods and data frequency. Concretely, we re-estimate our different models using two different data frequencies (daily and monthly) as well as a longer sample period from 1986 onwards.

The evidence globally indicates that our previous conclusions on the oil-stock return relationships hold for a longer sample period and regardless of data frequencies used. It means that empirical results we report in the paper are reasonably robust. The main differences we should mention are the following: *i)* when using monthly data, we find that the ARCH effects are slightly weaker and that the monthly frequency displays some bearing on asymmetry. *ii)* when daily data are used, AR(1) coefficients appear to be significant in higher number of cases, compared to weekly data.

#### *d) Real versus nominal prices*

The last step of our robustness check consists of examining whether our results subject to changes or not with respect to the effects of inflation. For this we repeat the same empirical analysis as above using real oil and stock prices instead of nominal prices. Globally, we realize that our results do not change significantly and our main conclusions remain valid, i.e., oil price changes positively affect sector returns in five cases (*Financials, Oil and Gas, Industrials, Basic Materials, and Consumer Services*) and negatively in three cases (*Food and Beverages, Health Care and Technology*). Weak negative relationships have been found for *Automobile and Parts* and *Personal and Household Goods* stock returns.

In summary, the evidence discussed in this section suggests that oil price changes significantly affect several European sectors, even after European market, world market and exchange rate risk factors as well as inflation effects are accounted for. These findings are also robust to the choice of the sample period and the frequency of data.

### **4. Some portfolio implications of the results**

In this section, we discuss some implications of the results we obtain for portfolio investment. First, we show that a model with oil risk presents superior out-of-sample forecasting results to a market model. Second, we illustrate how our results can be used in portfolio diversification and measure the out-of-sample benefit from portfolios considering these results. Note that we report, in what follows, the results obtained with weekly data over the out-of-sample sub-period running from November 20, 2008 to December 31, 2009 which totalizes 59 observations.<sup>9</sup>

#### **4.1 Out-of-sample forecasts of sector stock market returns**

Among the many issues involving portfolio investment and management, modelling and forecasting stock returns are one of the most intriguing topics that attract great interests from

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<sup>9</sup> The authors are very grateful to the anonymous referee and to Stéphane Grégoir for very interesting comments and valuable suggestions that have greatly motivated this part of the paper.

investors and researchers. Our analysis of causality and cycle synchronicity shows significant interactions between oil and sector stock returns as well as some predictability in stock price dynamics based on oil price changes. Moreover, as we have pointed out in sub-section 3.1, Models 2 and 3 provide better fit to our data than the market model (Model 1) in most cases. However, this does not guarantee that models with symmetric and asymmetric responses of stock returns to oil price changes will perform better in actual forecasting of stock returns.

Recall that the aim of forecasting evaluation is to minimize the expected loss, *i.e.* the difference between the predicted and actual returns. There is, up to date, a wide range of standard statistical loss functions that can be used to evaluate such a deviation in forecasting tasks. In this paper, we retain the most commonly used loss functions, namely Root Mean Squared Error (RMSE), Mean Absolute Error (MAE), and Mean of Absolute Percent Error (MAPE). They are defined as

$$RMSE_i = \sqrt{h^{-1} \sum_{t=T+1}^{T+h} (r_{it} - \hat{r}_{it})^2}$$

$$MAE_i = h^{-1} \sum_{t=T+1}^{T+h} |r_{it} - \hat{r}_{it}|$$

$$MAPE_i = h^{-1} \sum_{t=T+1}^{T+h} \left| \frac{r_{it} - \hat{r}_{it}}{r_{it}} \right|$$

where  $t$  denotes time period of the forecast sample,  $t = T+1, T+2, \dots, T+h$ .  $r_{it}$  and  $\hat{r}_{it}$  stand for the actual and forecasted returns respectively.

**Table 8. Forecasting evaluation results**

Sectors	Distribution	Model 1			Model 2			Model 3		
		RMSE	MAE	MAPE	RMSE	MAE	MAPE	RMSE	MAE	MAPE
Automobile and Parts	<i>Normal</i>	<b>0.0353</b>	<b>0.0267</b>	1.3141	0.0355	0.0270	<b>1.2726</b>	0.0356	0.0271	1.2764
	<i>Student-t</i>	<b>0.0348</b>	<b>0.0265</b>	1.2966	0.0353	0.0268	<b>1.2617</b>	0.0354	0.0268	1.2656
Financials	<i>Normal</i>	0.0293	0.0205	<b>0.7737</b>	<b>0.0286</b>	<b>0.0197</b>	0.7765	0.0288	0.0199	0.7768
	<i>Student-t</i>	0.0292	0.0207	<b>0.7765</b>	<b>0.0285</b>	<b>0.0194</b>	0.7770	0.0290	0.0201	0.7783
Food and Beverages	<i>Normal</i>	0.0195	0.0145	1.5296	0.0186	0.0138	<b>13797</b>	<b>0.0176</b>	<b>0.0133</b>	14006
	<i>Student-t</i>	0.0195	0.0146	1.5465	0.0184	0.0137	13846	<b>0.0172</b>	<b>0.0132</b>	<b>13831</b>
Oil and Gas	<i>Normal</i>	0.0195	0.0155	<b>0.9817</b>	<b>0.0185</b>	<b>0.0136</b>	1.2010	0.0189	0.0138	1.2173
	<i>Student-t</i>	0.0193	0.0154	<b>0.9905</b>	<b>0.0184</b>	<b>0.0133</b>	1.2169	0.0188	0.0138	1.2133
Health Care	<i>Normal</i>	0.0207	0.0155	1.8177	0.0198	0.0147	1.3644	<b>0.0188</b>	<b>0.0144</b>	<b>1.3595</b>
	<i>Student-t</i>	0.0207	0.0152	1.7528	0.0195	0.0142	<b>1.4761</b>	<b>0.0191</b>	<b>0.0141</b>	1.5023
Industrials	<i>Normal</i>	0.0147	0.0114	6.5534	<b>0.0143</b>	<b>0.0090</b>	<b>6.5283</b>	0.0148	0.0100	6.5977
	<i>Student-t</i>	0.0148	0.0116	6.6542	<b>0.0142</b>	<b>0.0091</b>	6.6051	0.0145	0.0102	<b>6.5091</b>
Basic Materials	<i>Normal</i>	0.0234	0.0195	0.8726	0.0216	<b>0.0178</b>	<b>0.8546</b>	<b>0.0212</b>	0.0186	0.8690
	<i>Student-t</i>	0.0223	0.0186	0.8955	0.0224	<b>0.0186</b>	<b>0.8463</b>	<b>0.0222</b>	0.0180	0.8607
Personal &	<i>Normal</i>	0.0136	0.0095	1.8676	0.0121	0.0088	0.9875	<b>0.0115</b>	<b>0.0076</b>	<b>0.9581</b>



Household Goods	<i>Student-t</i>	0.0128	0.0083	1.1230	0.0119	0.0089	0.9981	<b>0.0117</b>	<b>0.0080</b>	<b>0.9463</b>
Consumer Services	<i>Normal</i>	0.0145	0.0116	1.1535	<b>0.0136</b>	<b>0.0113</b>	<b>1.0934</b>	0.0139	0.0115	1.1098
	<i>Student-t</i>	0.0149	<b>0.0111</b>	1.1614	<b>0.0133</b>	0.0118	<b>1.1071</b>	0.0139	0.0117	1.0964
Technology	<i>Normal</i>	0.0285	<b>0.0200</b>	4.2735	0.0287	0.0205	4.0283	<b>0.0274</b>	0.0206	<b>3.6816</b>
	<i>Student-t</i>	<b>0.0272</b>	<b>0.0195</b>	4.0336	0.0277	0.0204	<b>2.6684</b>	0.0286	0.0204	3.8187
Telecommunications	<i>Normal</i>	0.0251	0.0177	5.4556	<b>0.0227</b>	0.0184	<b>4.0576</b>	0.0229	<b>0.0173</b>	7.0124
	<i>Student-t</i>	0.0256	0.0175	6.9991	<b>0.0221</b>	<b>0.0171</b>	<b>4.9122</b>	0.0225	0.0172	8.8588
Utilities	<i>Normal</i>	<b>0.0204</b>	<b>0.0165</b>	2.2235	0.0241	0.0180	<b>1.7849</b>	0.0209	0.0168	2.0726
	<i>Student-t</i>	<b>0.0203</b>	<b>0.0163</b>	2.2371	0.0220	0.0167	2.2030	0.0205	0.0164	<b>2.2005</b>

Notes: Bold numbers denote the lowest error statistics. Results are obtained with weekly data over the out-of-sample sub-period going from 11/20/2008 to 12/31/2009 (59 observations).

In Table 8, we provide out-of-sample forecast results for the three competitive models (*i.e.*, Models from 1 to 3). First, forecast evaluation functions show that the models estimated assuming the Student's *t*-distribution has in most cases the lowest losses and thus are superior to models assuming normal distribution. Second, Model 2 shows better forecasting results than Model 1 for the industries for which the in-sample estimations and evaluations indicate that oil price risk is significant. We thus conclude that augmenting the market model with oil price changes leads to better forecasting of sector stock returns in most cases. Finally, Model 3 shows superior results to Model 2 for the industries for which we have found significant asymmetries in the oil-stock returns relationships, *i.e.*, *Personal* and *Household Goods*, *Food* and *Beverages*, *Basic Materials*, and *Health Care*. As a result, for these industries taking into account asymmetric effects conducts to better stock returns forecasting and thus to better portfolio investment decisions.

#### **4.2 Result implications for out-of-sample portfolio diversification**

The main findings of our in-sample analysis can be summarized as follows: on average, stock and oil markets are weakly correlated, the reactions of stock returns to oil price changes differ greatly depending on the activity sector, and finally there are some signs of asymmetric interactions between oil and stock returns. In order to illustrate the out-of-sample benefit from portfolio diversification considering our results, we have adopted some index-based investment strategies by constructing several portfolios composed of both stocks and oil with different allocation rates, and report the main results of our portfolio simulations in Table 9.<sup>10</sup>

<sup>10</sup> Note that alternatively one can invest in futures crude oil contracts. However, since the correlations between returns on spot oil prices, and on 1-month and 3-month futures contracts are very high over our sample period (respectively 0.92 and 0.94), we have found similar results using futures oil contract returns in our portfolio simulations instead of oil price returns.

**Table 9. Stock portfolio diversification through Brent crude oil**

	Mean	Std. dev.	Sharpe ratio*	Sharpe ratio variation (%)	Sharpe ratio variation (%) (1)	Min.	Max.	Skew.	Kurt.
<i>Panel A</i>									
100% PF I	0.523	3.180	0.152	--	--	-8.808	3.180	-0.873	3.761
90% PF I	0.618	3.143	0.184	21.053	21.053	-8.558	6.605	-0.730	3.445
10% Crude Oil									
80% PF I	0.712	3.183	0.211	14.674	38.816	-8.357	8.096	-0.535	3.332
20% Crude Oil									
70% PF I	0.807	3.298	0.233	10.427	53.289	-8.131	9.586	-0.326	3.455
30% Crude Oil									
60% PF I	0.902	3.480	0.248	6.438	63.158	-7.905	11.077	-0.138	3.734
40% Crude Oil									
<i>Panel B</i>									
100% PF II	0.421	2.118	0.180	--	--	-5.743	4.619	-0.518	3.531
90% PF II	0.526	2.101	0.232	28.889	28.889	-4.902	5.192	-0.129	3.211
10% Crude Oil									
80% PF II	0.631	2.218	0.267	15.086	48.333	-4.054	6.840	0.224	3.491
20% Crude Oil									
70% PF II	0.736	2.451	0.284	6.367	57.778	-4.051	8.487	0.419	4.100
30% Crude Oil									
60% PF II	0.841	2.770	0.298	4.930	65.556	-4.882	10.135	0.474	4.600
40% Crude Oil									
<i>Panel C</i>									
100% PF III	0.789	4.008	0.187	--	--	-10.991	8.096	-0.789	3.676
90% PF III	0.875	3.902	0.214	14.439	14.439	-10.547	9.288	-0.623	3.553
10% Crude Oil									
80% PF III	0.925	3.857	0.230	7.477	22.995	-10.103	10.481	-0.428	3.517
20% Crude Oil									
70% PF III	0.994	3.876	0.247	7.391	32.086	-9.659	11.673	-0.230	3.608
30% Crude Oil									
60% PF III	1.064	3.958	0.259	4.858	38.503	-9.215	12.856	-0.052	3.813
40% Crude Oil									

Notes:

PF I is a weighted portfolio of all sectors, the DJ Stoxx portfolio;

PF II is an equally-weighted of sectors negatively linked to oil;

PF III is an equally-weighted of sectors positively linked to oil;

\*: Sharpe ratio computed using the European 3-month interest rate;

(1): Variation in Sharpe ratio compared with a portfolio of 100% stock (PF I in Panel A, PF II in Panel B, and PF III in Panel C).

Results are obtained with weekly data over the out-of-sample sub-period going from 11/20/2008 to 12/31/2009 (59 observations).

In Panel A of Table 9, we quantify the benefit of introducing crude Brent oil into a diversified portfolio of all industries. We show that adding oil into the DJ Stoxx portfolio permits to enhance its risk-return characteristics. For instance, with only 10% of the portfolio invested in Brent crude oil, the average weekly return increases from 0.523% to 0.618%, while the standard deviation decreases from 3.180% to 3.143%. Consequently, the Sharpe ratio moves from 0.152 to 0.184, and thus experiences an increase of more than 21%. The skewness and kurtosis fall from -0.873 to -0.730 and from 3.761 to 3.445 respectively. These results correspond effectively to our expectations given the weak correlations between oil price changes and sector stock returns discussed in Sections 2 and 3. However, the marginal benefit from including crude oil into the DJ Stoxx portfolio is reduced all the more so the

weight of crude oil in the combined portfolio increases as suggested by the Sharpe ratio's relative variations.

In Panels B and C, we construct two equally-weighted portfolios based on the results of this paper. The portfolio PF II is an equally-weighted portfolio containing all industries negatively linked to oil price changes, while the portfolio PF III is an equally-weighted portfolio containing all industries positively related to oil price changes. The inspection of the results in Panels B and C show that the inclusion of crude oil in PF II and PF III leads to a significant reduction of the portfolio's standard deviations while ensuring a higher average return. Accordingly, the Sharpe ratio increases substantially in both cases. The diversification gain is however two times larger for portfolio PF II than for portfolio PF III, which is not surprising given the results we established in Section 3. For instance, when 10% of the portfolio is dedicated to investments in Brent crude oil, the average weekly return in panel B moves up from 0.421% to 0.526%, and the standard deviation decreases from 2.118% to 2.101%, leading the Sharpe ratio grows by about 29%. At the same time, the investment risks associated with larger skewness and kurtosis fall as these coefficients diminish from -0.518 to only -0.129 and from 3.531 to 3.211 respectively. In Panel C, introducing 10% of oil into portfolio PF III leads to smaller improvements: the mean average increases from 0.789% to 0.875% and the standard deviation falls from 4.008% to 3.902%, and thus the Sharpe ratio increases by only 14.439%. The skewness and kurtosis decrease from -0.789 to only -0.623 and from 3.676 to 3.533, respectively.

Overall, our portfolio simulations, based on the paper's results, lead to conclude that investors have interests to keep an eye on crude oil market movements in their investment decision-making process. Further, they can improve the risk-adjusted return of their portfolios by simply adopting a sector-based portfolio investment strategy on the basis of this paper's results.

#### **4. Conclusion**

In this article, we investigated the linkages between oil and stock prices. Unlike other empirical investigations, which have focused largely on broad market indices (national and/or regional indices), we contribute to a better understanding of the relationship between oil prices and the stock markets in Europe by testing for short-term links at both the aggregate and sector by sector levels. Our results show strong significant linkages between oil price changes and stock markets for most European sectors. However, the nature and sensitivity of the reaction of stock returns to oil price shocks change considerably across sectors. Our out-

of-sample analysis reveals two important points. First, models with oil price changes as a factor of risk provide better forecasts of stock returns, and are thus more useful for portfolio investment decisions. Second, the inclusion of oil into a portfolio of stocks significantly improves its risk-return characteristics and a sector-based portfolio investment strategy considering the asymmetric responses of some industries to oil price shocks leads to a greater improvement of the Sharpe ratio.

Our findings should be of interest to researchers, regulators, and market participants. First, traders who are interested in investing in oil-sensitive stocks in Europe may, when oil prices are expected to increase, select stocks from sectors, such as *Oil* and *Gas*, with high positive sensitivity to oil prices. Alternatively, when oil price is expected to decrease, they may select sectors with negative sensitivity such as *Food* and *Beverages*. They can also use oil-related derivatives instruments. Thus, our results can be used to build profitable investment strategies. Second, the fact that sector returns in Europe have different sensitivities to oil price changes indicates valuable risk diversification opportunities across industries. Portfolio selection across rather than within sectors would be more efficient. Finally, investors and portfolio managers should rebalance their portfolios in keeping with their views of the sign of coming changes in oil prices (rises or falls), and our findings suggest that diversification can be achieved across sectors in all cases of oil price changes.

There are several avenues for future research. First, a sector analysis of the long-run linear and nonlinear links between oil and stock prices would be informative. Second, the econometric tools applied in this paper could be used to examine the effects of other energy products, such as natural gas. Third, a study of nonlinear causality between oil or other energy products and sector stock returns should be relevant. Finally, one of the future challenges would be to investigate whether oil price constitutes a common business cycle component across all European countries, that is affecting their sectoral indices indirectly.

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